

The prognosis of acute and persistent low-back pain: a meta-analysis

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ABSTRACT

Background: Although low-back pain is a highly prevalent condition, its clinical course remains uncertain. Our main objective was to systematically review the literature on the clinical course of pain and disability in patients with acute and persistent low-back pain. Our secondary objective was to investigate whether pain and disability have similar courses.

Methods: We performed a meta-analysis of inception cohort studies. We identified eligible studies by searching MEDLINE, Embase and CINAHL. We included prospective studies that enrolled an episode-inception cohort of patients with acute or persistent low-back pain and that measured pain, disability or recovery. Two independent reviewers extracted data and assessed methodologic quality. We used mixed models to determine pooled estimates of pain and disability over time.

Results: Data from 33 discrete cohorts (11 166 participants) were included in the review. The

variance-weighted mean pain score (out of a maximum score of 100) was 52 (95% CI 48–57) at baseline, 23 (95% CI 21–25) at 6 weeks, 12 (95% CI 9–15) at 26 weeks and 6 (95% CI 3–10) at 52 weeks after the onset of pain for cohorts with acute pain. Among cohorts with persistent pain, the variance-weighted mean pain score (out of 100) was 51 (95% CI 44–59) at baseline, 33 (95% CI 29–38) at 6 weeks, 26 (95% CI 20–33) at 26 weeks and 23 (95% CI 16–30) at 52 weeks after the onset of pain. The course of disability outcomes was similar to the time course of pain outcomes in the acute pain cohorts, but the pain outcomes were slightly worse than disability outcomes in the persistent pain cohorts.

Interpretation: Patients who presented with acute or persistent low-back pain improved markedly in the first six weeks. After that time improvement slowed. Low to moderate levels of pain and disability were still present at one year, especially in the cohorts with persistent pain.

Competing interests:

Christopher Maher has received grant funding from the National Health and Medical Research Council of Australia for the study of the prognosis of back pain. He has also received supplementary funding from GlaxoSmithKline for an investigator-initiated trial of paracetamol for acute low-back pain. James McAuley has received salary funding from the National Health and Medical Research Council of Australia and has received funding from the George Institute for Global Health for attending conferences. No competing interests declared for Luciola Costa, Robert Herbert, Mark Hancock and Leonardo Costa.

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Low-back pain is a highly prevalent condition associated with work absenteeism, disability and large health care costs; however, there is still disagreement about prognosis. For example, the European guidelines for the management of low-back pain states that 90% of patients with acute low-back pain recover in six weeks.¹ In contrast, some well-conducted cohort studies show a less optimistic picture, providing short-term estimates of recovery ranging from 39% to 76%.^{2,3} This wide range of estimates of prognosis is likely explained by differences in cohorts and definitions used to define the onset or conclusion of an episode of low-back pain. Because very different definitions of recovery are often used, it is difficult to obtain pooled estimates of recovery rates. Instead, it might be more useful to describe the clinical course of low-back pain in terms of expected changes in pain or disability over time.

A recent systematic review⁴ summarized the prognostic factors for persistent disabling low-back pain but did not describe the clinical course. The only meta-analysis to investigate the clinical course of acute low-back pain was published in 2003.⁵ This review concluded that both pain and disability improve rapidly within weeks (mean reduction of 58% of initial scores in the first month) and recurrences are common. A limitation of this review was that, although it retrieved 15 studies, only 5 were cohort studies; the remaining 10 were randomized controlled trials. Randomized trials often have narrow inclusion criteria and low rates of participation, which make them less suitable for inferring prognosis. The best design to describe the prognosis of a condition is a cohort study enrolling a representative sample of incident cases (i.e., by including patients at a similar early point in their condition).^{6,7} Such studies are known as inception

cohort studies. To the best of our knowledge, no review has yet investigated the clinical course of pain and disability among people with persistent low-back pain (subacute and chronic). Thus, the prognosis for people with persistent low-back pain is still uncertain.

The aim of our study was to systematically review the clinical course of pain and disability in patients with acute and persistent low-back pain. We included only inception cohort studies. Our second aim was to investigate whether pain and disability have similar courses.

Methods

Study selection

We searched MEDLINE, CINAHL and Embase databases from 1950 to Nov. 22, 2011, to identify potentially eligible studies. The search strategy used the study population terms suggested by the Cochrane Back Review Group together with a strategy for searching MEDLINE for prognosis studies. The terms used to identify the study population included “low-back pain,” “back pain” or “backache.” Terms used to identify prognosis studies included “inception,” “survival,” “life tables,” “log rank,” “prospective” and “follow-up studies.” The complete search strategies from all databases are included in Appendix 1 (available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.111271/-/DC1). Additionally, we performed a comprehensive examination of reference lists from eligible studies. We included only studies published in peer-reviewed journals. There were no language or geographic restrictions.

Two reviewers screened the titles and abstracts of each citation (Luciola Costa and Leo Costa). For each potentially eligible study, these authors obtained the full article and assessed whether the study fulfilled the inclusion criteria. In cases of disagreement, a third reviewer (C.M. or M.H.) was consulted and a decision was made by consensus. We defined low-back pain as pain or discomfort localized below the costal margins and above the inferior gluteal folds, with or without leg pain. Nonspecific low-back pain, the most common form of low-back pain, is a classification based on the exclusion of patients with a specific cause for their low-back pain (e.g., fracture, infection, cancer). It is sometimes called low-back pain of unknown origin.^{1,8}

We included studies that were prospective cohort studies and that included a well-defined inception cohort (episode inception) of participants with low-back pain. Specifically, we included studies that had (a) acute and subacute low-back pain cohorts that included participants with nonspecific low-back pain of less than 12

weeks duration and/or (b) persistent low-back pain cohorts that included participants with nonspecific low-back pain for more than 12 weeks but less than 12 months. For inclusion, the studies had to report pain or disability outcomes or a global measure of recovery.

We excluded studies that included patients with specific diseases such as tumours, fractures, arthritis or cauda equina syndrome (but not sciatica).

Data extraction

For each cohort, the following summary data were obtained: sample source, sample size, inception time, outcomes (pain, disability and recovery rates) and duration of follow-up. Continuous outcomes (pain and disability) were converted to a common 0–100 scale. Recovery rates were based on the definitions used in the individual articles. If there were insufficient data presented in a study report, we contacted the authors to request these data.

We classified the study cohorts as acute low-back pain cohorts or persistent low-back pain cohorts. This classification was based on the mean or median duration of pain at study entry. If the cohort had a mean or median duration of pain at study entry of less than six weeks, it was classified as an acute low-back pain cohort. If the study had mean or median duration of pain at study entry for greater than or equal to six weeks, the study cohort was classified as a persistent low-back pain cohort. When means or medians were not available, we used the midpoint of the range of the inclusion criteria for pain duration. We also extracted data with regard to prognostic factors from all cohorts. Because of the high variability of prognostic factors investigated by each study and the different statistical approaches used, we extracted data for univariate analyses only.

Quality assessment

We rated the quality of the studies using an adaptation of the methodologic criteria suggested by Altman.⁶ These criteria relate to sampling (2 items), completeness of follow-up (2 items) and description of prognostic outcomes (1 item). The same criteria have been used in previous systematic reviews on the prognosis of acute low-back pain⁵ and whiplash.⁹

Statistical analysis

We conducted a meta-analysis using the pain and disability outcome data. Pain and disability were modelled as a function of time. Time was treated in two ways. The first approach used the time of study entry as the reference time. We called this “inception time uncorrected.” The second

approach transformed the time data from each study by adding the mean or median inception time. We called this “inception time corrected.” Inception time uncorrected analyses provide prognoses that would apply at the time of study entry, whereas inception time corrected analyses provide prognoses that would apply at the time of onset of acute or persistent low-back pain.

We obtained pooled estimates of the time course of pain and disability using linear mixed models. The mixed models included a random study-specific intercept to account for the dependence of repeated observations within studies. Each study was assigned a weight equal to the inverse of the mean variance of the estimates from that study. We modelled time as a continuous variable. Both pain and disability declined linearly with log time, so all analyses used log

time as an independent variable. The reference time was considered to be day 1.

The regression model was used to obtain pooled point and interval estimates of outcomes at baseline and at 6, 26 and 52 weeks. The exception was that estimates of baseline values were not obtained for analyses of inception time corrected because this involved extrapolating the fitted curves substantially beyond the range of the data. To test if pain and disability had a similar clinical course, we combined the pain and disability data and added to the model an additional covariate coded 0 for pain outcomes or 1 for disability outcomes.

Results

We screened 28 613 potentially relevant studies

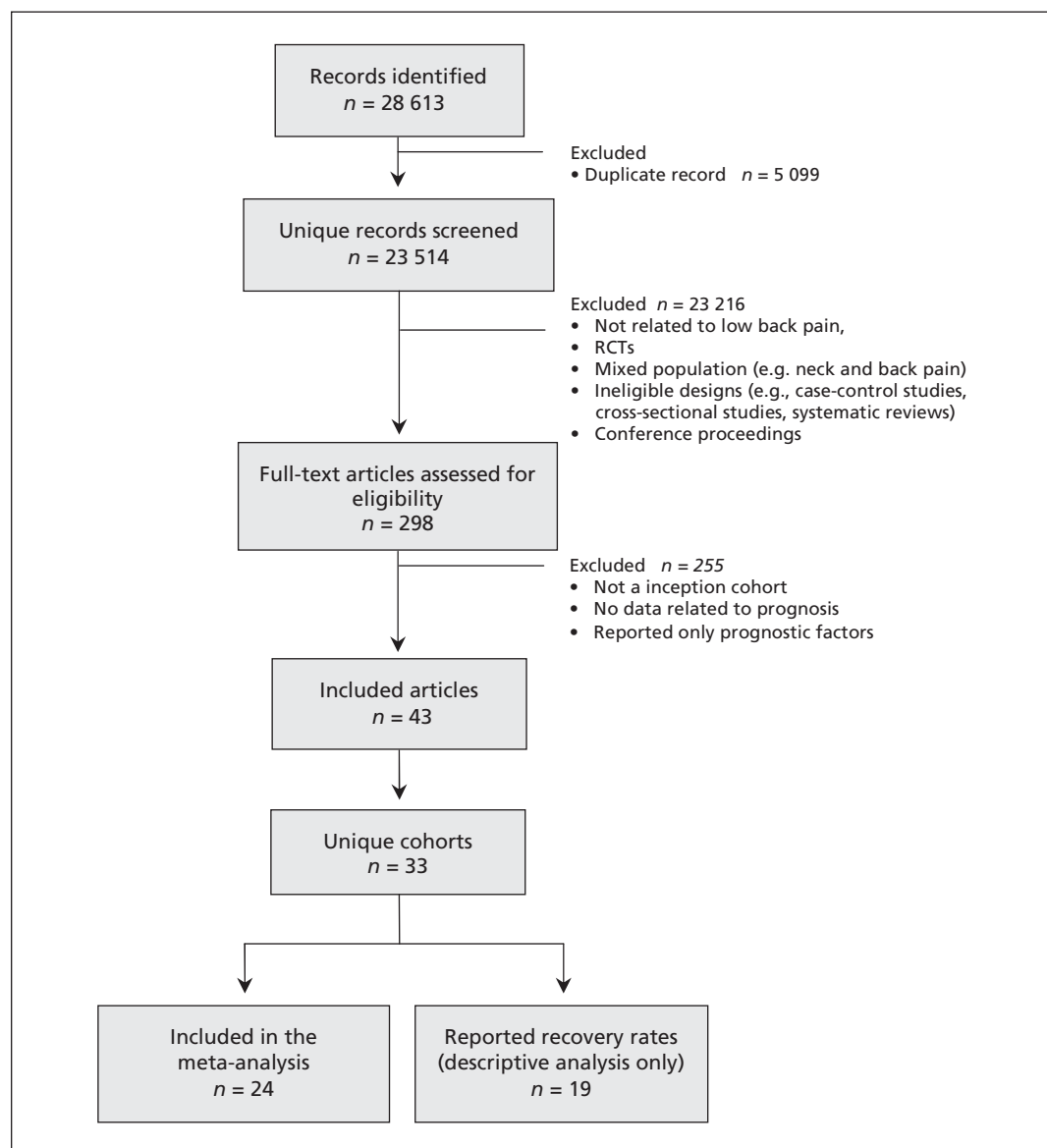


Figure 1: Selection of studies for inclusion in the systematic review. RCT = randomized controlled trials.

and evaluated 298 full-text articles. Of these, 43 articles reporting 33 cohorts (11 166 participants) met all criteria and were included in the review

(Figure 1). The description of the included cohorts is presented in Table 1. Many reports contained insufficient data; we contacted the authors of 24

Table 1 (part 1 of 2): Characteristics of the included studies

Study	Source of participants	Area	No. of participants	Onset of pain before study entry	Outcomes extracted	Follow-up
Bakker et al. ¹⁰	General practitioner	Netherlands	97	< 6 wk	Pain (NRS 0–10), recovery rates	6 mo
Bousema et al. ¹¹	General practitioner and advertisement	Netherlands	124	4–7 wk	Pain (VAS 0–100), disability (QBPDQ 0%–100%), recovery rates	12 mo
Breen et al. ¹²	Two primary care trust localities	England	75	< 12 wk	Pain (NRS 0–10), disability (RMDQ 0–24)	8 wk
Carey et al. ¹³ Sundararajan et al. ¹⁴	Urban and rural primary care and chiropractors, orthopedic and neurosurgeons, and health maintenance organization providers	US	1633	< 10 wk	Recovery rates	2, 4, 8, 12 and 24 wk
Carey et al. ¹⁵	Urban and rural primary care and chiropractors, orthopedic and neurosurgeons, health maintenance organization providers	US	96	12–22 wk	Disability (RMDQ 0–23), recovery rates	18 mo
Costa et al. ¹⁶	Primary care practitioners (general practitioners, physiotherapists and chiropractors)	Australia	406	12 wk	Pain (1–6), disability (1–5), recovery Rates	9 and 12 mo
Coste et al. ¹⁷	Primary care practitioners	France	103	< 72 hr	Pain (VAS 0–100); disability (RMDQ 0–24), recovery rates	8, 15 d and 1, 2 and 3 mo
Coste et al. ¹⁸	Primary care practitioners or rheumatologists	France	113	< 72 hr	Pain (VAS 0–100); disability (RMDQ 0–24), recovery rates	8, 15 d and 1, 2 and 3 mo
Epping-Jordan et al. ¹⁹ Shaw et al. ²⁰ Wahlgren et al. ²¹ Williams et al. ²²	Naval medical center	US	140	6–10 wk	Pain (DDS 0–20), disability (SIP 0%–100%), recovery rates	6 and 12 mo
Faber et al. ²³	General practitioners and occupational health physicians	Netherlands	103	3–12 wk	None	3 and 6 mo
Ferguson et al. ^{24,25}	Primary and urgent care facilities	US	32	< 4 wk	Pain (0–5), disability (VAS 0–150), recovery rates	2, 4, 6, 8, 10 wk and 6 mo
Gatchel et al. ^{26,27}	Industrial medicine clinic and orthopedic practices	US	421	< 6 wk	None	6 and 12 mo
Grotle et al. ^{2,28}	Primary health care providers	Norway	123	< 3 wk	Pain (NRS 0–10), disability (RMDQ 0–24), recovery rates.	1, 2, 3 and 4 wk and 3, 6, 9 and 12 mo
Gurcay et al. ²⁹	Tertiary care referral hospital	Turkey	99	< 3 wk	Recovery rates	1, 2, 4, 8 and 12 wk
Hasenbring et al. ³⁰	General or orthopedic practitioners	Germany	177	< 12 wk	Pain (NRS 0–10), disability (PDI 0–10)	6 mo

Continued

studies^{2,3,11-23,26-35,37-39,42-44,46,49} and received additional data for 18 cohorts.^{2,3,11-19,21,22,28,30,33,35,38,39,42-44,46,49} Our pooled analyses are based on summary data from 24 cohorts (4994 participants).

Methodologic quality

Almost all of the studies (97%) provided suffi-

cient definition of the sample: 17 (51.5%) studies explicitly described methods for assembling a representative sample. In total, 24 (72.7%) cohorts had a follow-up of at least 80%, and 87.9% of cohorts had a follow-up for at least one prognosis outcome at three months or longer (Table 2).

Table 1 (part 2 of 2): Characteristics of the included studies

Study	Source of participants	Area	No. of participants	Onset of pain before study entry	Outcomes extracted	Follow-up
Hazard et al. ³¹ Reid et al. ³²	Vermont Department of Labor and Industry	US	166	within 11 d of low-back pain	None	3 mo
Heneweer et al. ³³	Primary care physical therapy centres	Netherlands	66	< 12 wk	Pain (VAS 0–100), Disability (QBPDQ 0%–100%), recovery rates	2, 4, 8 and 12 wk
Henschke et al. ³	Primary care practitioners (general practitioners, physiotherapists and chiropractors)	Australia	969	> 24 h but < 2 wk	Pain (1–6), disability (1–5), recovery rates	6 wk, 3 and 12 mo
Klenerman et al. ³⁴	General practitioners	England	300	< 1 wk	Recovery rates	2 and 12 mo
Koleck et al. ³⁵	General practitioner	France	99	10–90 d	Pain (VAS 0–10), recovery rates	12 mo
Kovacs et al. ³⁶	Primary care centres	Spain	366	< 12 wk	Pain (VAS 0–10), disability (RMDQ 0–24)	2 wk and 8 wk
Lehmann et al. ³⁷	Occupational physicians	US	60	2–6 wk	None	6 mo
Melloh et al. ³⁸	Primary care settings	New Zealand	62	< 12 wk	Pain (NRS 0–100), disability (ODI 0–100), recovery rates,	3, 6, 12 wk and 6 mo
Poiraudeau et al. ³⁹	Rheumatologist	France	443	4–12 wk	Pain (4-points), disability (QBPDQ 0–20), recovery rates	3 mo
Reeser et al. ⁴⁰	Primary and tertiary care facilities	US	368	< 6 wk	Disability (MODEMS 0–100)	6 wk, 3 and 12 mo
Schiottz-Christensen et al. ⁴¹	General practices	Denmark	524	< 14 d	Recovery rates	1, 6 and 12 mo
Shaw et al. ⁴²⁻⁴⁴	Community-based occupational health clinics	US	568	< 2 wk	Pain (VAS 0–10), disability (RMDQ 0–100)	1 and 3 mo
Sieben et al. ⁴⁵	General practices	Netherlands and Belgium	44	< 2 wk	Disability (RMDQ 0–24)	2 wk, 3 and 12 mo
Sieben et al. ⁴⁶	General practices	Netherlands	222	< 3 wk	Pain (VAS 0–100), disability (QBPDQ 0–100)	3, 6 and 12 mo
Suri et al. ⁴⁷	Outpatient clinics	US	77	< 12 wk	Pain (NRS 0–10), disability (ODI 0–100)	6 wk
Swinkels-Meewisse et al. ⁴⁸	General practitioners and physiotherapists	Netherlands	555	< 4 wk	Pain (VAS 0–100), disability (RMDQ 0–24)	6 wk and 6 mo
Thomas et al. ⁴⁹	Community-based	US	43	< 4 wk	Pain (McGill), disability (RMDQ 0–24)	3, 6 and 12 wk
Valat et al. ⁵⁰	General practitioner and rheumatologist	France	2493	< 1 wk	Recovery rates	7 ± 1 wk of follow-up

Note: Note: DDS = Descriptor Differential Scale, MODEMS = Musculoskeletal Outcomes Data Evaluation and Management System, NRS = Numerical Rating Scale, ODI = Oswestry Disability Index, QBPDQ = Quebec Back Pain Disability Questionnaire, RMDQ = Roland Morris Disability Questionnaire, SIP = Sickness Impact Profile, VAS = Visual Analogue Scale.

Clinical course of pain and disability

Twenty-one cohorts reported pain outcomes. Fifteen cohorts^{2,3,10-12,17,18,24,25,28,30,36,42,43,44,46-49} were classified as acute low-back pain cohorts and six^{16,19-22,33,35,38,39} were persistent low-back pain cohorts. Twenty-two cohorts reported disability scores from at least one follow-up point: 16 of

these were acute low-back pain cohorts^{2,3,11,12,17,18,,24,25,28,30,36,40,42-49} and 6 were persistent low-back pain cohorts.^{15,16,19,20,21,22,33,38,39} Pain and disability data are shown in Figures 2 and 3. In the acute pain cohorts, the inception time uncorrected variance-weighted mean (95% confidence interval [CI]) pain score (out of a

Table 2: Methodologic quality assessment of included cohorts

Study	Defined sample*	Representative sample†	Follow-up rate > 80%‡	Follow-up time§	Prognosis¶
Bakker et al. ¹⁰	Yes	No	Yes	Yes	Yes
Bousema et al. ¹¹	Yes	No	Yes	Yes	Yes
Breen et al. ¹²	Yes	No	No	No	No
Carey et al. ¹³	Yes	Yes	Yes	Yes	Yes
Sundararajan et al. ¹⁴					
Carey et al. ¹⁵	Yes	Yes	No	Yes	Yes
Costa et al. ¹⁶	Yes	Yes	Yes	Yes	Yes
Coste et al. ¹⁷	Yes	Yes	Yes	Yes	Yes
Coste et al. ¹⁸	Yes	Yes	Yes	Yes	Yes
Epping-Jordan et al. ¹⁹	Yes	Yes	Yes	Yes	Yes
Shaw et al. ²⁰					
Wahlgren et al. ²¹					
Williams et al. ²²					
Faber et al. ²³	Yes	Yes	Yes	Yes	Yes
Ferguson et al. ^{24,25}	Yes	No	Yes	Yes	Yes
Gatchel et al. ^{26,27}	Yes	No	No	Yes	No
Grotle et al. ^{2,28}	Yes	No	Yes	Yes	Yes
Gurcay et al. ²⁹	Yes	Yes	Yes	Yes	Yes
Hasenbring et al. ³⁰	Yes	Yes	Yes	Yes	Yes
Hazard et al. ³¹	No	Yes	Yes	Yes	Yes
Reid et al. ³²					
Heneweer et al. ³³	Yes	Yes	Yes	Yes	Yes
Henschke et al. ³	Yes	Yes	Yes	Yes	Yes
Klenerman et al. ³⁴	Yes	No	No	Yes	Yes
Koleck et al. ³⁵	Yes	Yes	Yes	Yes	Yes
Kovacs et al. ³⁶	Yes	No	Yes	No	Yes
Lehmann et al. ³⁷	Yes	No	Yes	Yes	Yes
Melloh et al. ³⁸	Yes	Yes	Yes	Yes	Yes
Poiraudeau et al. ³⁹	Yes	Yes	Yes	Yes	Yes
Reeser et al. ⁴⁰	Yes	Yes	No	Yes	Yes
Schiottz-Christensen et al. ⁴¹	Yes	No	Yes	Yes	Yes
Shaw et al. ⁴²⁻⁴⁴	Yes	No	Yes	Yes	Yes
Sieben et al. ⁴⁵	Yes	No	No	Yes	Yes
Sieben et al. ⁴⁶	Yes	No	No	Yes	Yes
Suri et al. ⁴⁷	Yes	Yes	No	No	Yes
Swinkels-Meewisse et al. ⁴⁸	Yes	No	No	Yes	Yes
Thomas et al. ⁴⁹	Yes	No	Yes	Yes	No
Valat et al. ⁵⁰	Yes	No	Yes	No	Yes

*Description of source of patients and inclusion and exclusion criteria.
†Participants were selected as consecutive cases.
‡Outcome data were available for at least 80% of participants at one follow-up point.
§At least one prognostic outcome was followed up at three months or later.
¶Raw data, percentages, survival rates or continuous outcome reported.

maximum of 100) was 52 (95% CI 48–57) at baseline, 23 (95% CI 21–25) at 6 weeks, 12 (95% CI 9–15) at 26 weeks and 6 (95% CI 3–10) at 52 weeks after entering the study. In the cohort with persistent pain, the score was 51 (95% CI 44–59) at baseline, 33 (95% CI 29–38) at 6 weeks, 26 (95% CI 20–33) at 26 weeks and 23 (95% CI 16–30) at 52 weeks. The estimates of pain and disability (inception time corrected and uncorrected) in the acute and persistent low-back pain cohorts are shown in Table 3.

Within-study variation

The standard deviation of pain or disability outcomes was typically about 20 points on a 100-point scale. This was true for pain and disability outcomes, for acute and persistent back pain and for short-term and longer-term outcomes. This is indicative of a moderate degree of person-to-person variability in outcomes.

Between-study variation

We used two statistics to quantify between-study variability. The first was tau, which is a model-derived estimate of the standard deviation of outcomes on day 1. The second measure of between-study variability was rho, the proportion of the total variance due to between-study variation on day 1. For the inception time uncorrected analyses, between-study variability was moderate: tau ranged from 8 to 14 and rho ranged from 0.60 to 0.73. For the inception time corrected analyses, between-study heterogeneity was moderate to high, particularly for the disability outcome: tau ranged from 14 to 22 and rho ranged from 0.65 to 0.97.

Comparison of the clinical course of pain and disability

The clinical course of pain and disability in the acute low-back pain cohorts was similar and the

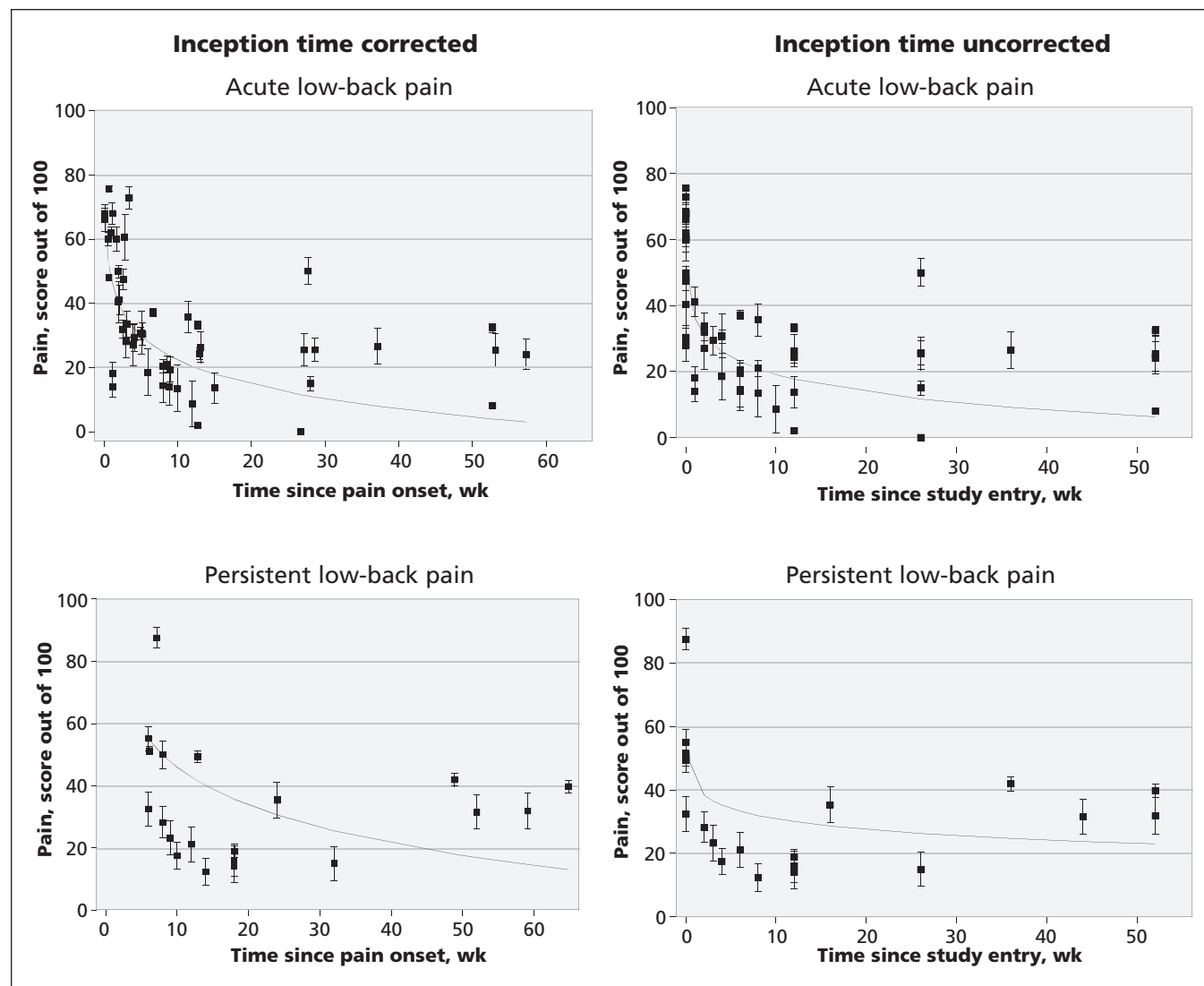


Figure 2: Course of pain in patients with acute and chronic low-back pain. Cohorts are represented by individual data points.

difference was not statistically significant ($p = 0.2$). However, the clinical course of pain was more favourable than the clinical course of disability in the persistent pain cohorts ($p = 0.002$).

Recovery from low-back pain

Recovery was measured in 19 (57.6%) of the included cohorts. Given the heterogeneity of the studies in terms of the definitions of low-back pain, recovery and length of follow-up, it was not possible to pool the data. Most studies reported that the majority of patients with acute low-back pain had recovered by 12 weeks. Among patients with persistent low-back pain, recovery occurred in less than half of the patients at the longest follow-up time. Recovery rate data are presented in Table 4.

Prognostic factors

Appendix 2 (available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.111271/-/DC1) presents a description of prognostic factors evaluated by each

study (when available). A large number of potential prognostic factors were tested, with demographic and clinical characteristics being the most frequently evaluated. Reporting of the strength of association was very inconsistent among the studies, which made it impossible to pool these data.

Interpretation

The results of our meta-analysis of the course of acute and persistent low-back pain based on data from inception cohort studies are consistent with findings from previous studies of acute low-back pain. Our review confirms the broad finding of previous reviews that the typical course of acute low-back pain is initially favourable: there is a marked reduction in mean pain and disability in the first six weeks. Beyond six weeks, improvement slows and thereafter only small reductions in mean pain and disability are apparent up to one year. By one year, the average levels of pain and disability for acute low-back pain were low (mean pain score [inception time uncorrected] of

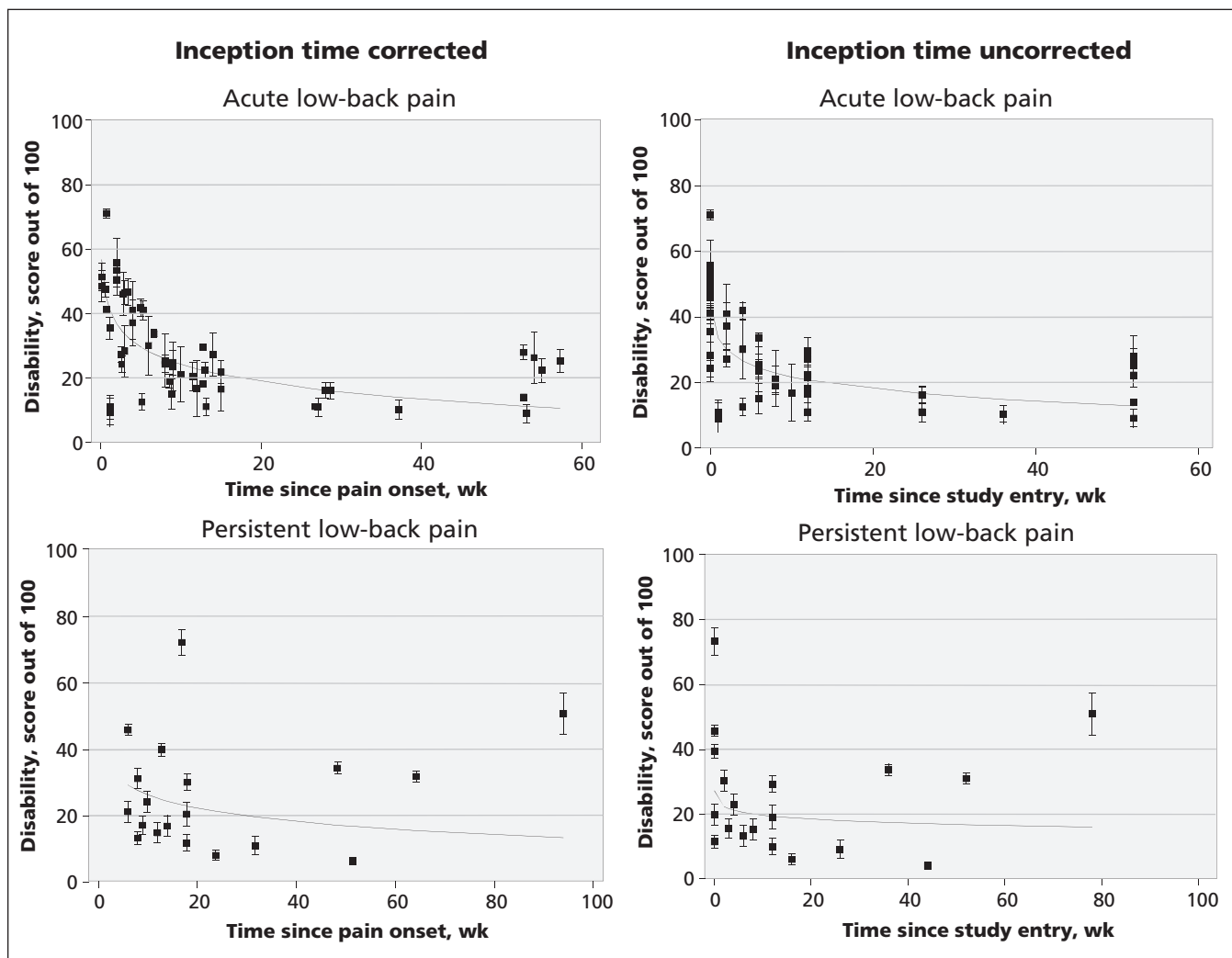


Figure 3: Course of disability in patients with acute and chronic low-back pain. Cohorts are represented by individual data points.

6 and disability score of 13), suggesting that patients can expect to have minimal pain or disability at one year.

People with persistent low-back pain also experienced substantial improvement in the first six weeks, but there were only very small reductions in average pain and disability between 6 and 52 weeks (mean reductions in pain and disability score [inception time uncorrected] of 10 and 4, respectively). Patients with persistent low-back pain could expect to have moderate levels of pain and disability at 12 months (mean pain and disability scores of 23 and 17, respectively). For both acute and persistent low-back pain, there was moderate between-participant variability in outcomes.

We used a sensitive search strategy, which necessitated screening a large number of titles (28 613 titles screened, 43 eligible articles, number needed to screen = 665 titles). This means that it is likely that most eligible studies were included in our review. We included only inception cohort studies because this type of design minimizes bias in studies of prognosis.⁷ We were able to pool pain and disability outcomes from most of the eligible cohorts. Unfortunately, six studies did not report continuous measures of pain or disability, and the authors of these studies did not respond to requests to provide data. Therefore, not all studies that collected continuous measures of pain and disability were included in our final analysis.

In this study, pooled estimates of the clinical course of pain and disability were determined using mixed linear models. There are at least two reasons why these data may provide inaccurate prognoses. First, there is obvious between-study

heterogeneity. That is, the mean outcomes of some cohorts differ substantially from the pooled estimate. The cause of this heterogeneity is not known, but it could be real (e.g., because of variation in characteristics of particular study cohorts) or due to bias (e.g., high rates of loss to follow-up in some cohorts). A second issue is that there was evidence of a moderate amount of person-to-person variability within each study (which could be due to, for example, initial pain intensity, duration of symptoms, previous episodes, presence or absence of sciatica, personal characteristics and care received). The standard deviations for pain and disability were consistently around 20. As the outcomes were normally distributed, about one-third of patients would have prognoses that differ by more than 20 points from the study mean. The best way to explore this variability and to generate subject-specific prognoses that take into account suspected prognostic factors would be to perform a meta-analysis using individual patient data. It may be possible to conduct such studies in the future.

Another possible explanation for the improvement of patients over time may be regression to the mean, which is a consequence of random variation over time. Some of the randomness will be “true” variation caused by biological, psychological or sociological mechanisms (e.g., people’s pain will fluctuate unpredictably because of their activity levels and mood) and some will be measurement error (because people’s self-reports of their pain and disability will not be perfectly precise). It is probably impossible to properly partition out the contributions of these two sources of regression to the mean. But

Table 3: Pooled estimates of mean pain and disability*

Group	Pooled estimate of mean pain (95% CI)		Pooled estimate of mean disability (95% CI)	
	Inception time uncorrected	Inception time corrected	Inception time uncorrected	Inception time corrected
Acute low-back pain				
Baseline	52 (48–57)	69 (61–78)	45 (42–48)	57 (52–62)
6 wk	23 (21–25)	28 (25–31)	24 (23–26)	28 (26–30)
26 wk	12 (9–15)	12 (8–15)	16 (15–18)	17 (15–19)
52 wk	6 (3–10)	4 (0–9)	13 (11–15)	11 (9–14)
Persistent low-back pain				
Baseline	51 (44–59)	NA†	27 (24–30)	51 (39–63)
6 wk	33 (29–38)	55 (46–63)	21 (19–22)	28 (25–31)
26 wk	26 (20–33)	29 (23–35)	18 (16–20)	19 (18–21)
52 wk	23 (16–30)	17 (7–27)	17 (14–19)	15 (12–18)

Note: CI = confidence interval, NA = not available.
 *All values are expressed in a scale ranging from 0 (i.e., no pain or disability) to 100 (i.e., maximum pain or disability).
 †The estimates of baseline values were not obtained for analyses of inception time corrected time for pain outcomes because this involved extrapolating the fitted curves substantially beyond the range of the data.

Table 4: Summary of recovery from low back pain

Study	Inception time	Recovery rates
Bakker et al. ¹⁰	< 6 wk	<ul style="list-style-type: none"> 40% fully recovered by 12 wk and did not experience a recurrence within 6 mo
Bousema et al. ¹¹	4–7 wk	<ul style="list-style-type: none"> After 1 yr, 32% had no back pain complaints
Carey et al. ¹³ Sundararajan et al. ¹⁴	< 10 wk	<ul style="list-style-type: none"> 95% functionally recovered (a return to a functional status similar to that before the onset of low-back pain) by 6 mo 69% considered themselves completely better at 6 mo
Carey ¹⁵	12–22 weeks	<ul style="list-style-type: none"> 33% had no functionally limiting symptoms after 18 mo 16% had no back symptoms after 18 mo
Costa ¹⁶	12 weeks	<ul style="list-style-type: none"> Only 11% of patients had not returned to work in their previous capacity at the onset of chronicity and, of those 46%, had returned to work by 12 mo The cumulative probability of being pain-free, maintained for one mo, was 35% at 9 mo and 42% at 12 mo after onset The cumulative probability of having no disability, maintained for one mo, was 39% at 9 mo and 47% at 12 mo after onset The cumulative probability of being completely recovered (pain-free, no disability and returned to prior work status maintained for one mo) was 35% at 9 mo and 41% at 12 mo after onset of chronic pain
Coste et al. ¹⁷	< 72 h	<ul style="list-style-type: none"> 90% (95%CI 84%–96%) of patients recovered within the first 2 wk 98% (95%CI 95%–100%) of patients recovered after 3 mo 40% of patients lost no time from work and return to work was slower than recovery from back pain
Coste et al. ¹⁸	< 72 h	<ul style="list-style-type: none"> 87% (95%CI 79%–95%) had recovered in 30 d 95% (95%CI 91%–100%) had recovered in 3 mo
Epping-Jordan et al. ¹⁹ Shaw et al. ²⁰ Wahlgren et al. ²¹ Williams et al. ²²	6–10 wk	<ul style="list-style-type: none"> 54% improved by 6 mo 67% improved by 12 mo
Ferguson et al. ^{24,25}	<4 wk	<ul style="list-style-type: none"> 68% of patients were not impaired according to functional performance, and 80% were not impaired in terms of pain by 14 to 18 wk
Grotle et al. ^{2,28}	< 3 wk	<ul style="list-style-type: none"> 76% had recovered (RMDQ < 4) after both 4 wk and 3 mo 83% of patients had recovered fully from their disability after 1 yr (RMDQ < 4)
Gurcay et al. ²⁹	< 3 wk	<ul style="list-style-type: none"> 27% recovered in the first wk 31% recovered in the second wk 23% recovered in the fourth wk 10% recovered at the eighth wk 1% recovered after 12 wk 9% developed chronic low-back pain
Heneweer et al. ³³	< 12 wk	<ul style="list-style-type: none"> 52% recovered in less than 4 wk 55% recovered at 12 wk, of whom 76% did not report work absenteeism
Henschke et al. ³	> 24 h to < 2 wk	<ul style="list-style-type: none"> The cumulative probability of returning to work with pre-back pain work status and duties for those who reduced their work status at baseline because of low-back pain was 80% at 6 wk, 83% at 12 wk and 90% by 1 yr The cumulative probability of having no disability was 55% at 6 wk, 73% by 12 wk and 83% by 1 yr The cumulative probability of being pain free was 39% by 6 wk, 58% by 12 wk and 73% by 1 yr The cumulative probability of being completely recovered was 39% by 6 wk, 57% by 12 wk and 72% by 1 yr
Klenerman et al. ³⁴	< 1 wk	<ul style="list-style-type: none"> 21% were classified as having no pain at 12 mo 28% were classified as having no intermittent pain at 12 mo 93% were classified as not having constant pain at 12 mo
Koleck et al. ³⁵	10–90 d	<ul style="list-style-type: none"> 67% of patients were classified as improved after 1 yr 67% were classified as not having chronic low-back pain after 1 yr
Melloh et al. ³⁸	< 12 wk	<ul style="list-style-type: none"> 75% were classified as nonpersistent low-back pain at 6 wk 25% were classified as persistent low-back pain at 6 wk
Poiraudeau et al. ³⁹	4–12 wk	<ul style="list-style-type: none"> 60% did not have persistent low-back pain at 3 mo 59% did not use sick leave during the 3 mo period 83% had returned to work at 3 mo
Schiottz-Christensen et al. ⁴¹	< 14 d	<ul style="list-style-type: none"> 84% (95%CI 80%–87%) of patients functionally recovered at 6 mo 92% (95%CI 89%–94%) functionally recovered at 12 mo 47% of patients completely recovered at 6 and 12 mo
Valat et al. ⁵⁰	< 1 wk	<ul style="list-style-type: none"> 61.3% had fully recovered at 7 wk

Note: CI = confidence interval, RMDQ = Roland Morris Disability Questionnaire.

for practical purposes, it is probably not necessary. The data have a useful pragmatic interpretation because they indicate the time course of self-reported reductions in pain and disability, even though the cause of the reduction is not known.

To our knowledge, there has been only one previous systematic review that has quantitatively synthesized study data on the clinical course of low-back pain.⁵ A strength of our review is that we retrieved more studies than the previous review, and we used an analysis approach that could accommodate outcomes measured at different times. We were thus able to provide more precise estimates of the course of acute low-back pain. The availability of more cohorts also made it practical to not include data from randomized controlled trials. Typically, the inclusion and exclusion criteria in randomized controlled trials are more restrictive than those used in cohort studies,⁷ which means that cohort studies are likely to provide more generalizable estimates of prognosis. Unlike the previous quantitative synthesis that only considered the clinical course of acute low-back pain,⁵ we also examined the clinical course of persistent low-back pain.

Although we acknowledge that the identification of prognostic factors is important, it is a challenge to summarize this information because there is great variability in the methods used to measure putative prognostic factors, build prognostic models, and quantify and report prognostic value.^{6,51} These difficulties were also observed in previous systematic reviews of acute low-back pain^{4,5} and whiplash.⁹ One systematic review on the course of acute low-back pain identified distress, previous episodes of low-back pain and job satisfaction as likely prognostic factors.⁵ The most recent review of prognostic factors for the development of persistent low-back pain⁴ found that maladaptive pain coping behaviours, nonorganic signs, functional impairment, general health status and presence of psychiatric comorbidities were the most important prognostic factors.

Limitations

Only a small proportion of eligible studies explicitly excluded patients with leg pain or sciatica,^{17,18} while most studies either did not provide any information on this distinction (12 cohorts) or included some patients with leg pain or sciatica (19 cohorts). It was common for authors to refer to the presence of leg pain as sciatica without differentiating whether the pain was somatic or radicular. Therefore, it is possible that our estimates of the prognosis of nonspecific low-back pain may be influenced by the inclusion of some patients with radicular pain.

We classified studies as acute or persistent based on the median or mean duration of symptoms when provided or the midpoint of the range if not available. It is possible that some studies with means or medians close to our threshold of six weeks could be similar to other cohorts with a different classification. Only one study³⁹ that provided accurate information on the duration of symptoms had a mean duration within one week of our threshold. Studies including patients typically considered to have subacute low-back pain (6–12 weeks) were included in our persistent group. Although it is possible that these patients have a more favourable prognosis, the reasonably consistent pattern of results in our persistent cohorts suggests this was not the case.

Conclusion

We found that patients with acute or persistent low-back pain improved markedly in the first six weeks, but beyond this time improvement slowed. Even at one year, patients had low to moderate levels of pain and disability. Future research is needed to make more precise recommendations on the standardization of recovery definitions of a low-back pain episode and to develop an optimal search strategy to retrieve prognostic studies. These improvements in research will facilitate the development of future prognostic systematic reviews.

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