A cute ischemic stroke is usually classified as thrombosis or embolism. However, acute cerebral hypoperfusion can also precipitate a specific type of ischemic stroke called hemodynamic stroke. Alpha-blockers are strong vasodilators originally classified as antihypertensive agents. Their effectiveness in relieving symptoms in the lower urinary tract has led to their use for the treatment of benign prostatic hyperplasia since 1976. The Antihypertensive and Lipid-Lowering Treatment to Prevent Heart Attack Trial (ALLHAT) showed that α-blockers were inferior to other classes of drugs as the first-line therapy for hypertension. Currently, α-blockers are most commonly prescribed for the treatment of benign prostatic hyperplasia. Nevertheless, they have inherent adverse cardiovascular effects, including dizziness and hypotension, especially soon after initiation (known as the first-dose effect), because of their effects on lowering blood pressure. Researchers have reported an increase in the risk of hypotension-related adverse events, including hypotension, syncope and fractures, during the initiation period of α-blocker therapy. Our research group found that the use of α-blockers was associated with an increased risk of hip and femur fractures during the early initiation period among older men without hypertension. Whether the acute hypotensive effect of α-blockers can precipitate the episode of ischemic stroke has not been clarified in a large-scale study, although a case of acute hemiparesis after one dose of doxazosin has been reported.

The aim of this study was to provide a quantitative estimate of the risk of ischemic stroke during the early initiation period of α-blocker therapy among older men. We used a self-controlled case series design based on claims data from Taiwan’s National Health Insurance claims database.

Methods

Data source

The National Health Insurance Research claims database has been described previously. To
Comply with Taiwanese privacy regulations, all personal identifiers are encrypted and data can be analyzed only anonymously. The protocol for this study was approved by the Institutional Review Board of the National Taiwan University Hospital, which waived requirement for informed consent.

**Study design and cohort definition**

We used claims data from the National Health Insurance Research database for the period 2006–2009 and applied a self-controlled case series design. We identified all men aged 50 years or more as of Jan. 1, 2007, who had continuous coverage under the National Health Insurance program for at least 12 months. The study period for each case was defined as the period from Jan. 1, 2007, to Dec. 31, 2009. We selected men who were incident users of α-blockers and had a diagnosis of ischemic stroke within the 3-year study period. We excluded those who had any diagnosis concerning ischemic stroke or any prescription of α-blocker in the baseline period (Jan. 1, 2006, to Dec. 31, 2006). Only the first event of ischemic stroke was included in the analysis. Because all patients with ischemic stroke were identified from the inpatient records, we presumed that the dates of hospital admission were the dates of occurrence of ischemic stroke.

**Definitions of exposures and events**

We identified the use of 4 α-blockers: terazosin, doxazosin, tamsulosin and alfuzosin. We defined ischemic stroke according to the presence of International Classification of Diseases, Ninth Revision, Clinical Modification (ICD-9-CM) codes 433.x1, 434.x1, 435.9, 436, 437.1x or 437.9x in the discharge diagnoses in the inpatient records of the National Health Insurance Research claims database. The diagnosis of acute ischemic stroke in the database has been validated recently.

**Study timeframe**

In the database’s outpatient records, each drug dispensed at outpatient clinics or pharmacies is included, with the corresponding date of dispensing. All drugs prescribed during hospital stays, however, are recorded in the database’s inpatient records, with only the dates of hospital admission and discharge. Therefore, we defined the index date either as the dispensing date of α-blocker therapy identified from outpatient records or as the date of hospital admission if the first dose of α-blocker was prescribed during the hospital stay. Because all patients with ischemic stroke were identified from the inpatient records, we presumed that the dates of hospital admission were the dates of occurrence of ischemic stroke.

We operationally defined the early initiation period (≤ 21 d after the index date) as post-exposure risk period 1 and the 22–60 days after the index date as post-exposure risk period 2. To determine whether cases of exposure were event-dependent, we also defined the 21 days before the index date as pre-exposure risk period 1 and the 22–60 days before the index date as pre-exposure risk period 2. The times before and after the pre- and post-exposure risk periods in the 3-year study period were defined as the unexposed periods (Figure 1). For analysis, we grouped the 2 unexposed periods into 1 unexposed period because they were far away from the index date, they conveyed no acute hypotensive effect of α-blockers, and the crude incidence rates of ischemic stroke in the 2 unexposed periods were similar (see Table S1 in Appendix 1, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.150624/-/DC1). Because we could not determine the true temporal relationship between initial exposure and the event for cases occurring on the same day, we excluded the index date during partition of the risk periods and included it in the unexposed period in the primary analysis for all the participants to get more conservative estimates.

**Baseline characteristics**

We evaluated the baseline comorbidities of the study population using the Elixhauser Comorbidity Index and coded them as binary vari-

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**Figure 1:** Pictorial representation of the study design.
ables. We determined the specific comorbidity as the presence of corresponding diagnoses in at least 2 outpatient visits or at least 1 inpatient record within the 12-month period before the index date. Only items with a prevalence of more than 2% were retained.

**Statistical analysis**

Using the conditional Poisson regression model, we compared the incidence rates of ischemic stroke in various risk periods with the rates in the unexposed period and presented the results as incidence rate ratios (IRRs).\(^\text{10,13,14,17,18}\) The key advantage of a self-controlled case series design is that time-invariant confounders, such as sex, location, genetics and underlying state of health, are controlled for implicitly.\(^\text{14}\) In the conditional Poisson regression model, we adjusted only for age at the index date and development of atrial flutter or fibrillation (ICD-9-CM code 427.3), which was included as a time-varying term.

Because the acute hypotension-related adverse events associated with α-blockers may be modified by the status of concomitant antihypertensive treatment,\(^\text{9,10}\) we repeated the primary analysis with stratification according to whether patients were taking concomitant antihypertensive therapy on the index date. We also performed a stratification analysis according to different α-blockers administered.

We performed 2 sensitivity analyses to test the robustness of our results. In the first analysis, we included the index date in pre-exposure risk period 1, on the assumption that all patients whose index date and stroke event occurred on the same day had their α-blocker therapy prescribed during the hospital stay after a preceding ischemic stroke. In the second analysis, we excluded all patients who began their α-blocker therapy during a hospital admission (regardless of whether it was the same admission for stroke).

We repeated the same design as the primary analysis for patients newly prescribed diuretics and calcium-channel blockers, to explore whether the observed results for α-blockers represented a possible group effect of antihypertensive agents.

We performed all analyses using SAS software, version 9.2 (SAS Institute Inc.).

**Results**

We identified a total of 7502 patients who fulfilled our study criteria (1663 were prescribed terazoin, 2252 doxazosin [1493 standard formula and 759 modified-release formula], 3108 tamsulosin and 479 alfuzosin). The mean age was 71.0 years at the index date of α-blocker therapy and at the first occurrence of ischemic stroke (Table S2 in Appendix 1). Most of the ischemic strokes occurred within the 3-month period around the index date (Figure S1 in Appendix 1).

For the whole study population, the incidence of ischemic stroke was significantly higher in post-exposure risk period 1 than in the unexposed period (adjusted IRR 1.40, 95% confidence interval [CI] 1.22–1.61). It was also significantly higher in both pre-exposure risk periods (pre-exposure risk period 1: adjusted IRR 2.87, 95% CI 2.60–3.17; pre-exposure risk period 2: adjusted IRR 1.99, 95% CI 1.82–2.18). Conversely, the incidence in post-exposure risk period 2 was lower than in the unexposed period (adjusted IRR 0.80, 95% CI 0.70–0.92) (Table 1).

Among patients with concomitant prescriptions for antihypertensive agents, the incidence of ischemic stroke in post-exposure risk period 1 was similar to that in the unexposed period.

<table>
<thead>
<tr>
<th>Table 1: Relative risk of ischemic stroke before and after the start of α-blocker therapy (index date) among men aged 50 years or older</th>
</tr>
</thead>
<tbody>
<tr>
<td>Period*</td>
</tr>
<tr>
<td>---</td>
</tr>
<tr>
<td>Unexposed (ref)</td>
</tr>
<tr>
<td>Pre-exposure</td>
</tr>
<tr>
<td>Risk period 2 (−22 to −60 d)</td>
</tr>
<tr>
<td>Risk period 1 (−1 to −21 d)</td>
</tr>
<tr>
<td>Post-exposure</td>
</tr>
<tr>
<td>Risk period 1 (≤ 21 d)</td>
</tr>
<tr>
<td>Risk period 2 (22 to 60 d)</td>
</tr>
</tbody>
</table>

Note: CI = confidence interval, IRR = incidence rate ratio, ref = reference category.

*See the Methods and Figure 1 for details about the periods.
†Adjusted for age and development of atrial flutter/fibrillation.
(adjusted IRR 1.07, 95% CI 0.88–1.29), whereas the incidence in post-exposure risk period 2 was significantly lower than in the unexposed period (adjusted IRR 0.67, 95% CI 0.56–0.81). In contrast, among patients without concomitant prescriptions for antihypertensive agents, the incidence of ischemic stroke was significantly higher in post-exposure risk period 1 than in the unexposed period (adjusted IRR 2.11, 95% CI 1.73–2.57), whereas the incidence in post-exposure risk period 2 was similar to that in the unexposed period (adjusted IRR 1.06, 95% CI 0.86–1.31).

Irrespective of whether patients had concomitant prescriptions for antihypertensive agents, the incidence of ischemic stroke was significantly higher in both pre-exposure risk periods than in the unexposed period (Table 2).

Regardless of the type of α-blocker administered, the adjusted IRRs of ischemic stroke were significantly elevated in both pre-exposure risk periods compared with the unexposed period (Table 3). In post-exposure risk period 1, the adjusted IRR remained significantly elevated for all of the α-blockers studied except alfuzosin. In post-exposure risk period 2, none of the α-blockers was associated with an increased risk of ischemic stroke, and alruzosin and the standard formula of doxazosin were associated with a decreased incidence of ischemic stroke (Table 3).

In the first sensitivity analysis, the ischemic strokes that occurred on the same day as the index date of α-blocker therapy (n = 1793; 23.9% of the study population) were included in pre-exposure risk period 1. In the second sensitivity analysis, the ischemic strokes that occurred on the same day as the index date of α-blocker therapy were included in pre-exposure risk period 2. The adjusted IRRs of ischemic stroke were significantly elevated in both pre-exposure risk periods compared with the unexposed period (Table 3). In post-exposure risk period 1, the adjusted IRR remained significantly elevated for all of the α-blockers studied except alfuzosin. In post-exposure risk period 2, none of the α-blockers was associated with an increased risk of ischemic stroke, and alruzosin and the standard formula of doxazosin were associated with a decreased incidence of ischemic stroke (Table 3).

### Table 2: Relative risk of ischemic stroke among patients with and without concomitant exposure to other antihypertensive agents at index date of α-blocker therapy

<table>
<thead>
<tr>
<th>Period*</th>
<th>Concomitant use of antihypertensive agents</th>
<th>No concomitant use of antihypertensive agents</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Unadjusted IRR (95% CI)</td>
<td>Adjusted IRR† (95% CI)</td>
</tr>
<tr>
<td>Unexposed (ref)</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Pre-exposure</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk period 2 (−22 to −60 d)</td>
<td>2.01 (1.80–2.24)</td>
<td>2.05 (1.84–2.28)</td>
</tr>
<tr>
<td>Risk period 1 (−1 to −21 d)</td>
<td>2.91 (2.58–3.28)</td>
<td>2.94 (2.61–3.32)</td>
</tr>
<tr>
<td>Post-exposure</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk period 1 (≤ 21 d)</td>
<td>1.09 (0.90–1.32)</td>
<td>1.07 (0.88–1.29)</td>
</tr>
<tr>
<td>Risk period 2 (22 to 60 d)</td>
<td>0.69 (0.58–0.83)</td>
<td>0.67 (0.56–0.81)</td>
</tr>
</tbody>
</table>

Note: CI = confidence interval, IRR = incidence rate ratio, ref = reference category.
*See the Methods and Figure 1 for details about the periods.
†Adjusted for age and development of atrial flutter/fibrillation.

### Table 3: Relative risk of ischemic stroke, by α-blocker

<table>
<thead>
<tr>
<th>Period*</th>
<th>Terazosin</th>
<th>Doxazosin</th>
<th>Doxazosin-XL</th>
<th>Tamsulosin</th>
<th>Alfuzosin</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Unadjusted IRR (95% CI)</td>
<td>Adjusted IRR† (95% CI)</td>
<td>Unadjusted IRR (95% CI)</td>
<td>Adjusted IRR† (95% CI)</td>
<td>Unadjusted IRR (95% CI)</td>
</tr>
<tr>
<td>Unexposed (ref)</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Pre-exposure</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk period 2 (−22 to −60 d)</td>
<td>1.77 (1.45–2.15)</td>
<td>1.80 (1.48–2.20)</td>
<td>1.92 (1.57–2.35)</td>
<td>1.95 (1.59–2.38)</td>
<td>2.56 (1.98–3.31)</td>
</tr>
<tr>
<td>Risk period 1 (−1 to −21 d)</td>
<td>2.80 (2.27–3.45)</td>
<td>2.85 (2.31–3.52)</td>
<td>2.55 (2.03–3.20)</td>
<td>2.57 (2.04–3.23)</td>
<td>3.36 (2.49–4.52)</td>
</tr>
<tr>
<td>Post-exposure</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk period 1 (≤ 21 d)</td>
<td>1.37 (1.02–1.83)</td>
<td>1.37 (1.02–1.83)</td>
<td>1.57 (1.17–2.09)</td>
<td>1.53 (1.15–2.05)</td>
<td>1.61 (1.06–2.45)</td>
</tr>
<tr>
<td>Risk period 2 (22 to 60 d)</td>
<td>0.79 (0.59–1.06)</td>
<td>0.79 (0.59–1.05)</td>
<td>0.65 (0.46–0.91)</td>
<td>0.64 (0.46–0.89)</td>
<td>1.12 (0.76–1.63)</td>
</tr>
</tbody>
</table>

Note: CI = confidence interval, IRR = incidence rate ratio, ref = reference category, XL = modified-release formula.
*See the Methods and Figure 1 for details about the periods.
†Adjusting for age and development of atrial flutter/fibrillation.
tivity analysis, we excluded the patients whose α-blocker therapy was started during a hospital stay (n = 3135; 41.8% of study population). Owing to the reduction in event numbers in the unexposed period compared with the design of the primary analysis, all of the IRRs of ischemic stroke in the 4 risk periods compared with the unexposed period were increased in both sensitivity analyses compared with the primary analysis (Tables S3–S6 in Appendix 1).

When we repeated the same design as the primary analysis among patients newly prescribed diuretics, we observed results similar to those of the primary analysis (Table S7 in Appendix 1). Among patients newly prescribed calcium-channel blockers, the risk of ischemic stroke was increased in post-exposure risk period 1 regardless of whether the patients had concomitant prescriptions for antihypertensive agents (Table S8 in Appendix 1).

**Interpretation**

In our study of claims data for more than 7500 older men in Taiwan prescribed an α-blocker, we observed an increased risk of ischemic stroke in the early initiation period (within 21 d after the start of treatment) and a reduced risk in the late initiation period (22–60 d after the start of treatment). In our stratified analysis, only patients without concomitant exposure to other antihypertensive agents had an increased risk of ischemic stroke during the early initiation period of α-blocker therapy. In contrast, patients already using antihypertensive agents had no increased risk of ischemic stroke in the early post-exposure period and even had a decreased risk of stroke after treatment for more than 3 weeks. The increased risk of ischemic stroke during the early initiation period of α-blocker therapy was observed for nearly all of the α-blockers studied.

Alpha-blockers are notorious for their first-dose effect of acute hypotension during the early initiation period. Although several studies have shown that α-blocker therapy could improve cerebral blood flow after treatment for 4–8 weeks, acute hypotension associated with antihypertensive therapy could precipitate cerebral hypoperfusion and ischemic stroke.

One possible explanation of our findings was that patients without underlying hypertension were vulnerable to the first-dose effect of α-blockers and contributed to the observed increase in risk of ischemic stroke in the early initiation period of α-blockers in the whole study population. In contrast, patients with underlying hypertension appeared to be tolerant of the first-dose effect of α-blockers, benefited from the positive influence of α-blockers on cerebral blood flow and contributed to the observed decrease in risk of ischemic stroke in the late initiation period in the whole study population.

Significantly high IRRs in the 2 pre-exposure risk periods reflected the prescription pattern of α-blocker therapy frequently being started after an ischemic stroke (Table S9 in Appendix 1). We could attribute this prescription pattern to the frequent development of high blood pressure during the episode of an acute ischemic stroke. However, this explanation could not be applied to tamulosin and alfuzosin, the 2 α-blockers without an indication for use in patients with hypertension. Therefore, development of urinary symptoms after an ischemic stroke could be a better explanation of this prescription pattern, since most α-blockers are currently prescribed for benign prostatic hyperplasia rather than hypertension.

It has been reported that patients with stenosis of the intracranial carotid artery are more prone to ischemic stroke provoked by cerebral hypoperfusion than patients with stenosis of the extracranial carotid artery. Chinese people are at increased risk of intracranial carotid artery stenosis, and white people are at increased risk of extracranial carotid artery stenosis. Our findings were derived from claims data for mainly ethnic Chinese men in Taiwan. Further research is needed to determine whether the same findings would be observed in other ethnic populations. In addition, studies investigating the value of transcranial Doppler examination in the diagnosis of intracranial carotid artery stenosis and guidance of α-blocker therapy are warranted.

In the ALLHAT study, patients in the doxazosin group had a slightly increased rate of stroke compared with those in the chlorthalidone group (relative risk 1.19, 95% CI 1.01–1.40; p = 0.04). In our study, the IRR of ischemic stroke during the early initiation period among patients without underlying hypertension was higher among those prescribed α-blockers than among those prescribed diuretics (adjusted IRR 2.11 v. 1.80); therefore, our findings are in line with those from the ALLHAT study despite indirect comparison. Among patients prescribed calcium-channel blockers, the IRR of ischemic stroke was much higher in the early initiation period than in the unexposed period regardless of concomitant use of other antihypertensive agents. More pharmacoepidemiologic studies are needed before the group effect of all hypertensive agents on acute cerebral hypoperfusion and acute ischemic stroke can be established.

**Limitations**

We did not include 5-α-reductase inhibitors, an alternative treatment for benign prostatic hyperplasia, in our analysis owing to an inadequate
sample size. Another limitation was our inability to determine from the claims data whether prescribed medications were taken by the patients.

**Conclusion**

The association between α-blocker therapy and ischemic stroke was driven mostly by the prescription of α-blockers following a previous ischemic stroke. However, our results presented evidence that α-blockers were associated with an increase in the risk of ischemic stroke during the early initiation period, especially among patients who were not taking other antihypertensive agents. We recommend caution when prescribing α-blockers to patients who are not taking other antihypertensive medications.

**References**


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**Contributors:** All of the authors contributed to the study conception and design and the data interpretation. Raymond Kuo and Ho-Min Chen conducted the statistical analysis. Chao-Lun Lai and Arnold Chan drafted the manuscript. All of the authors revised the manuscript for important intellectual content, approved the final version to be published and agreed to act as guarantors of the work.

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