

# Impact of pharmacist administration of influenza vaccines on uptake in Canada

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## ABSTRACT

**BACKGROUND:** Uptake of influenza vaccination in Canada remains suboptimal despite widespread public funding. To increase access, several provinces have implemented policies permitting pharmacists to administer influenza vaccines in community pharmacies. We examined the impact of such policies on the uptake of seasonal influenza vaccination in Canada.

**METHODS:** We pooled data from the 2007–2014 cycles of the Canadian Community Health Survey ( $n = 481\,526$ ). To

determine the impact of influenza vaccine administration by pharmacists, we estimated the prevalence ratio for the association between the presence of a pharmacist policy and individual-level vaccine uptake using a modified Poisson regression model (dependent variable: vaccine uptake) with normalized weights while controlling for numerous health and sociodemographic factors.

**RESULTS:** Across all survey cycles combined, 28.8% of respondents reported

receiving a seasonal influenza vaccine during the 12 months before survey participation. Introduction of a policy for pharmacist administration of influenza vaccine was associated with a modest increase in coverage (2.2%) and an individual's likelihood of uptake (adjusted prevalence ratio 1.05, 95% confidence interval 1.02–1.08).

**INTERPRETATION:** Uptake of influenza immunization was modestly increased in Canadian jurisdictions that allowed pharmacists to administer influenza vaccines.

Despite widespread public funding for influenza vaccination in Canada, either for high-risk groups or the entire population depending on the province or territory, uptake remains suboptimal.<sup>1</sup> Numerous factors influence an individual's decision to be vaccinated, and although many of these relate to health beliefs and risk–benefit perceptions that can be challenging to address, ease of access to vaccination may play a role.<sup>2–5</sup>

Influenza vaccines are primarily administered by nurses and physicians in Canada, but several provinces have recently implemented policies permitting administration by pharmacists (Table 1).<sup>6</sup> Community pharmacists who have been authorized to administer influenza vaccines could help to overcome issues with accessibility given their ubiquitous distribution, extended working hours, walk-in policies and availability to people without a primary care provider.<sup>7–9</sup>

In the United States, all states allow trained pharmacists to administer vaccines.<sup>10</sup> This has enhanced patient access as perceived by patients,<sup>11</sup> elicited support from prospective vaccinees<sup>12</sup> and increased vaccine coverage.<sup>13,14</sup> Steyer and colleagues<sup>14</sup> com-

pared influenza vaccine coverage in the US from 1995 to 1999 and found an absolute increase of 10.7% among adults aged 65 years and older in states where pharmacists could give the vaccine (from 57.7% to 68.4%), as compared with an increase of 3.5% in states that did not have such a policy at the time of the study (from 61.2% to 64.7%). Grabenstein and colleagues<sup>13</sup> found a significant increase in influenza vaccination among younger adults. Improvements in influenza vaccine coverage consistent with a pharmacist vaccinator effect have also been observed in rural populations in the US<sup>15</sup> and in populations in Japan.<sup>16</sup>

In Canada, patients have expressed support for pharmacist-administered vaccination,<sup>17</sup> and pharmacists have indicated strong interest in expanding their scope of practice.<sup>18</sup> A Canadian pilot study suggested that 80% of influenza vaccine recipients prefer getting their vaccine at pharmacy-based clinics.<sup>7</sup> However, it is unknown whether these findings translate to population-based increases in vaccine coverage, particularly given changes over time in other factors such as removal of financial barriers to influenza immunization through public funding. We sought to de-

termine the impact of regulatory changes allowing pharmacists to provide influenza vaccines on the uptake of seasonal influenza immunization in Canada.

## Methods

### Study population

We used nationally representative data from the 2007–2014 cycles of the Canadian Community Health Survey. This cross-sectional survey has been conducted by Statistics Canada through telephone and in-person interviews annually since 2007, covering a range of questions related to health status, health care utilization and determinants of health.<sup>19</sup> Using a multistage stratified cluster design, the survey includes a sample of about 65 000 respondents aged 12 years and older in each cycle. The survey excludes people living on reserves and other Aboriginal settlements, full-time members of the Canadian military, people in institutions and residents in remote health regions in Quebec (all representing < 3% of the population).<sup>19</sup> The response rate declined across survey cycles, from 77.6% to 65.6%.<sup>20</sup>

Our study was approved by the Ethics Review Board of Public Health Ontario.

### Definitions

The dependent variable was self-reported influenza vaccination, determined through responses to the questions “Have you ever had a (seasonal) flu shot?” and “When did you have your last (seasonal) flu shot?” Individuals who reported receiving their last flu shot in the 12 months before the survey were considered immu-

nized. These respondents were also asked, “In which month did you have your last flu shot?” Those whose response matched the month of the survey date were then asked, “Was that this year or last year?” We considered respondents who reported receiving the vaccine during the same month as the survey date but in the preceding year as not immunized.

We constructed our primary independent variable — the ability for pharmacists to administer publicly funded influenza vaccines (“pharmacist policy”) — based on province of residence and influenza season (Table 1). The start of the pharmacist policy was deemed to coincide with the start of the corresponding influenza season, defined as Oct. 1 to the following Sept. 30. Residents of Alberta and British Columbia from the 2009/10 influenza season onward, New Brunswick from 2010/11 onward, Ontario from 2012/13 onward and Nova Scotia from 2013/14 onward were classified as living in a jurisdiction with a pharmacist policy. Although Alberta introduced its policy in 2007, few pharmacists received certification for the 2007/08 and 2008/09 influenza seasons, and pharmacists were not part of the publicly funded program until 2009/10. The available survey data do not cover subsequent influenza seasons when Manitoba, Newfoundland and Labrador, Prince Edward Island and Saskatchewan introduced their pharmacist policies.

We chose the following potential confounders for influenza vaccination a priori based on a review of the existing literature: age, sex, rural residence (population concentration < 1000 and population density < 400 per square kilometer), household income, education, marital status, immigration status, racial/cultural group, self-reported health status, presence of chronic conditions (asthma, chronic obstructive pulmonary disease, heart disease, stroke, diabetes, cancer), body mass index, having a child less than 5 years old in the household, smoking status, receipt of the A/H1N1 2009 pandemic influenza vaccine (only for those surveyed in the 2010 cycle) and having a regular doctor.<sup>21–23</sup> We also examined the impact of living in a jurisdiction that provides universal publicly funded influenza vaccination (based on province and influenza season) as a separate policy intervention from the pharmacist policies.<sup>21</sup> To account for underlying temporal and provincial variations, we included influenza season and province as potential confounders.

### Statistical analysis

We pooled individual-level responses from all survey cycles, with the final data set considered to represent the characteristics of the average population over the combined period of the survey cycles.<sup>24</sup> Normalized weights were used to account for an unequal probability of selection in the sample. We used the Cochran–Armitage test to assess temporal trends in influenza vaccination by province. We used a weighted modified Poisson regression model to estimate the prevalence ratio for the association between the presence of a pharmacist policy and individual-level influenza vaccine uptake,<sup>25</sup> controlling for sociodemographic information, health status and health behaviours associated with influenza vaccination. We chose this approach over logistic regression analysis because of the frequency of the outcome.<sup>26</sup> We chose Quebec as the reference jurisdiction because it had the lowest ob-

**Table 1: Timing of implementation of universal funding for influenza vaccination and policy permitting pharmacists to administer publicly funded influenza vaccines, by province/territory**

Province/territory	Universal funding*	Pharmacist policy
Newfoundland and Labrador	2014	2014
Prince Edward Island	2004	2014
Nova Scotia†	2010	2013
New Brunswick†	NA	2010
Quebec	NA	NA
Ontario†	2000	2012
Manitoba	2010	2014
Saskatchewan	2010	2015
Alberta†	2009	2009
British Columbia†	NA	2009
Nunavut	2005	NA
Northwest Territories	2003	NA
Yukon	1999	NA

Note: NA = policy not available.

\*Influenza vaccines are freely available for residents aged 6 mo and older; in the Yukon, the policy applies to residents 18 yr and older.

†Province with pharmacist policy in survey cycles analyzed.

Table 2: Characteristics of the study population\*

Characteristic	Total sample		Province or influenza season with pharmacist policy		Province or influenza season without pharmacist policy	
	No. (%)	Influenza vaccination, %	No. (%)	Influenza vaccination, %	No. (%)	Influenza vaccination, %
Weighted sample	481 526 (100)	28.8	134 250 (27.9)	30.4	347 276 (72.1)	28.2
Age group, yr						
12–19	52 287 (10.9)	20.3	14 078 (10.5)	22.6	38 209 (11.0)	19.4
20–49	237 901 (49.4)	18.5	66 510 (49.5)	19.9	171 391 (49.4)	18.0
50–64	116 846 (24.3)	32.4	32 933 (24.5)	34.4	83 913 (24.2)	31.7
≥ 65	74 492 (15.5)	61.8	20 729 (15.4)	62.9	53 763 (15.5)	61.3
Sex						
Male	235 935 (49.0)	25.7	66 019 (49.2)	27.1	169 916 (48.9)	25.1
Female	245 591 (51.0)	31.7	68 231 (50.8)	33.5	177 360 (51.1)	31.1
Province/territory						
Newfoundland and Labrador	7366 (1.5)	24.3	NA	NA	7366 (2.1)	24.3
Prince Edward Island	2039 (0.4)	29.7	NA	NA	2039 (0.6)	29.7
Nova Scotia	13 411 (2.8)	42.1	1777 (1.3)	44.7	11 634 (3.4)	41.7
New Brunswick	10 596 (2.2)	32.7	5515 (4.1)	35.6	5081 (1.5)	29.6
Quebec	112 554 (23.4)	22.5	NA	NA	112 554 (32.4)	22.5
Ontario	188 273 (39.1)	31.6	50 328 (37.5)	31.3	137 945 (39.7)	31.6
Manitoba	16 309 (3.4)	27.3	NA	NA	16 309 (4.7)	27.3
Saskatchewan	13 793 (2.9)	28.3	NA	NA	13 793 (4.0)	28.3
Alberta	51 815 (10.8)	28.1	34 505 (25.7)	28.5	17 310 (5.0)	27.4
British Columbia	64 006 (13.3)	29.6	42 125 (31.4)	29.5	21 881 (6.3)	29.8
Nunavut	307 (0.1)	34.9	NA	NA	307 (0.1)	34.9
Northwest Territories	574 (0.1)	33.5	NA	NA	574 (0.2)	33.5
Yukon	483 (0.1)	30.7	NA	NA	483 (0.1)	30.7
Chronic condition						
Asthma	39 498 (8.2)	36.2	10 624 (7.9)	36.8	28 874 (8.3)	35.9
COPD	10 504 (2.2)	55.2	3395 (2.5)	54.9	7109 (2.0)	55.4
Heart disease	22 336 (4.6)	58.2	5519 (4.1)	61.6	16 817 (4.8)	57.1
Stroke	4596 (1.0)	53.3	1260 (0.9)	54.4	3336 (1.0)	52.9
Diabetes	29 421 (6.1)	55.4	7934 (5.9)	57.6	21 487 (6.2)	54.6
Cancer	9109 (1.9)	54.3	2421 (1.8)	54.1	6688 (1.9)	54.4
Has regular doctor						
Yes	407 405 (84.6)	31.7	116 191 (86.5)	33.0	291 214 (83.9)	31.2
No	73 611 (15.3)	12.6	17 907 (13.3)	13.6	55 704 (16.0)	12.3
Influenza season						
2006/07	44 966 (9.3)	31.8	NA	NA	44 966 (12.9)	31.8
2007/08	60 294 (12.5)	29.6	NA	NA	60 294 (17.4)	29.6
2008/09	60 788 (12.6)	29.9	NA	NA	60 788 (17.5)	29.9
2009/10	61 070 (12.7)	25.6	14 742 (11.0)	28.9	46 328 (13.3)	24.5
2010/11	62 221 (12.9)	28.3	16 254 (12.1)	29.4	45 967 (13.2)	27.9
2011/12	62 857 (13.1)	27.7	16 698 (12.4)	29.1	46 159 (13.3)	27.2
2012/13	64 668 (13.4)	27.6	42 223 (31.5)	29.1	22 445 (6.5)	24.7
2013/14	64 662 (13.4)	30.6	44 333 (33.0)	32.9	20 329 (5.9)	25.5
Universal funding for influenza vaccines						
Yes	249 139 (51.7)	31.2	86 609 (64.5)	30.5	162 530 (46.8)	31.7
No	232 387 (48.3)	26.1	47 641 (35.5)	30.2	184 746 (53.2)	25.1

Note: COPD = chronic obstructive pulmonary disease, NA = pharmacist policy did not exist in province or influenza season studied.

\*Additional characteristics of the study population are available in Appendix 1 (available at [www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1](http://www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1)).

served levels of influenza vaccination (thus all prevalence ratios comparing provinces and territories would be greater than 1) and because it was the only province that had not implemented either universal funding or a pharmacist policy (thus it was akin to a “no exposure” group). We assessed the model for interactions of pharmacist policy with age, chronic conditions and income.

We performed a sensitivity analysis in which we set introduction of the pharmacist policy in Alberta to the 2007/08 and 2008/09 seasons, as opposed to the 2009/10 season.

A significance level of  $p < 0.05$  was used for all tests. We used SAS statistical software (version 9.3, SAS Institute Incorporated) for all analyses.

## Results

After we excluded respondents with missing data on influenza vaccination ( $n = 15\,764$ ), we had a total weighted sample of 481 526 for our analysis (Table 2, and Appendix 1, available at [www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1](http://www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1)). Across all survey cycles, 28.8% of respondents reported receiving a seasonal influenza vaccine in the 12 months before survey participation. The proportion was slightly higher in provinces with a pharmacist policy than in those without a policy (30.4% v. 28.2%). The proportion of respondents who reported receiving an influenza vaccine decreased over the study period nationally ( $p < 0.001$ ). Decreases over time were observed in Quebec and Ontario, whereas increases were observed in Nova Scotia, New Brunswick and Alberta (Table 3). We observed notable season-to-season variability in the temporal trends.

Influenza vaccination remained higher in provinces with a pharmacist policy than in those without a policy after we adjusted for potential confounders (adjusted prevalence ratio 1.05, 95%

confidence interval [CI] 1.02–1.08) (Table 4). We observed no significant interaction of pharmacist policies with age, chronic conditions or income. Respondents in provinces with universal funding for influenza vaccination were more likely than those in provinces without a universal funding policy to report receiving seasonal influenza vaccines (adjusted prevalence ratio 1.13, 95% CI 1.10–1.17), independent of the presence of a pharmacist policy.

Influenza vaccination uptake was positively associated with increased age ( $\geq 50$  yr), female sex, urban residence, high household income, high educational attainment, being married, being of Asian descent, having chronic conditions (except for post-stroke effects), having a higher body mass index, having a child less than 5 years of age in the household, receiving the A/H1N1 2009 pandemic influenza and having a regular doctor (Table 4 and Appendix 2, available at [www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1](http://www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1)). Uptake was negatively associated with being an immigrant, reporting better health status and daily smoking. Influenza vaccination varied considerably across seasons and provinces.

In the sensitivity analysis, we found no change in the impact of pharmacist policies when we used the 2007/08 or 2008/09 influenza season as the first year of the policy for Alberta.

## Interpretation

During the study period, influenza vaccine coverage declined over time in Canada, but there was substantial interprovincial variability in this trend. Individuals living in provinces with a policy allowing administration of publicly funded influenza vaccines by pharmacists were more likely to report receipt of a seasonal influenza vaccine in the year before survey participation than those living in jurisdictions without a pharmacist policy. Based on the available data capturing the early implementation period of pharmacy poli-

**Table 3: Proportion of participants who reported receipt of influenza vaccine within 12 months before survey, by province/territory and influenza season**

Province/territory	2006/07	2007/08	2008/09	2009/10	2010/11	2011/12	2012/13	2013/14	<i>p</i> value	Direction of trend
Newfoundland and Labrador	22.4	23.4	24.7	23.4	24.7	24.1	24.4	26.5	0.08	–
Prince Edward Island	32.8	28.0	26.6	26.2	30.3	32.0	27.7	34.9	0.2	–
Nova Scotia	40.0	39.7	37.0	43.4	47.4	43.3	40.4	44.7*	< 0.001	Increase
New Brunswick	28.5	29.7	28.0	32.0	36.4*	34.2*	34.8*	37.2*	< 0.001	Increase
Quebec	25.1	25.1	24.9	16.4	21.1	21.6	22.4	23.9	< 0.001	Decrease
Ontario	37.5	33.6	33.5	26.9	30.6	29.4	29.7*	32.9*	< 0.001	Decrease
Manitoba	27.6	25.1	27.3	30.5	25.9	26.1	26.7	29.6	0.2	–
Saskatchewan	26.8	28.1	28.4	26.5	30.2	30.5	25.2	30.4	0.2	–
Alberta	28.0	25.7	28.6	28.5*	29.3*	26.8*	27.2*	30.6*	< 0.001	Increase
British Columbia	32.1	28.4	29.4	29.2*	28.3*	30.2*	28.2*	31.7*	0.7	–
Nunavut	36.6	34.3	44.0	41.8	43.1	26.0	26.8	33.4	0.2	–
Northwest Territories	35.1	32.8	29.6	33.4	39.0	35.9	29.0	34.7	0.9	–
Yukon	29.2	27.4	23.6	45.3	32.7	26.8	26.8	30.0	0.8	–

\*Pharmacist policy in place.

cies for most provinces, the impact of the policies so far has been modest.

Pharmacist policies could conceivably increase influenza vaccination through 3 mechanisms related to accessibility: improved availability, geographic proximity and accommodation.<sup>27</sup> Large numbers of community pharmacists are available to administer vaccines; for example, 51% of eligible pharmacists in Ontario ( $n = 7358$ ) have been authorized and trained to do so as of 2014.<sup>28</sup> In addition, most Canadians live close to a pharmacy, thereby possibly increasing physical access to these services. For example, 79% of the population in Nova Scotia and 91% of the population in Ontario live within a 5-km driving distance of a community pharmacy.<sup>29,30</sup> This is particularly important in rural areas, where access to medical or public health services is often limited.<sup>9,31</sup> Community pharmacies are also accommodating: people can visit during longer hours of operation, often without appointments, and generally experience shorter wait times.<sup>7,9,32,33</sup> A study in the United Kingdom found that 51% of patients surveyed who were eligible for free influenza vaccination through a physician preferred instead to pay for the service at a pharmacy because of convenience, and 43% chose to pay because of ease of access.<sup>34</sup>

Although the results of our study suggest a small impact associated with pharmacist policies, at least during the initial years of implementation, there may be other potential benefits associated with these policies. For instance, enhanced advertising of seasonal influenza vaccination by pharmacies may increase awareness of the important burden of influenza and promote other health-promoting behaviours such as hand-washing and respiratory etiquette, even if it does not necessarily translate to increased influenza vaccine uptake. Because 80% of Canadians consult with pharmacists,<sup>35</sup> those who seek nonvaccination services from pharmacists may also receive reminders for annual influenza vaccination. This may be an effective strategy to increase uptake in certain high-risk groups who are less likely to get the vaccine, such as smokers. However, to achieve the intended policy objective of increased vaccine coverage in the population, more promotion of influenza vaccination and further facilitation of vaccine delivery (e.g., vaccinating children at school) may be required in addition to pharmacist policies.

### Strengths and limitations

Our study is novel in assessing the impact of pharmacist policies on influenza vaccination coverage across all Canadian jurisdictions over time. The inclusion of multiple cycles of the Canadian Community Health Survey resulted in a large sample, which allowed us to incorporate many important covariates known to

**Table 4: Unadjusted and adjusted prevalence ratios for influenza vaccination**

Variable	Prevalence ratio (95%CI)	
	Unadjusted	Adjusted*
Presence of pharmacist policy	1.08 (1.06–1.10)	1.05 (1.02–1.08)
Universal funding for influenza vaccines	1.20 (1.18–1.21)	1.13 (1.10–1.17)
Age group, yr		
12–19 (ref)	1.00 (ref)	1.00 (ref)
20–49	0.91 (0.88–0.94)	0.78 (0.74–0.81)
50–64	1.60 (1.55–1.65)	1.18 (1.12–1.24)
≥ 65	3.05 (2.96–3.13)	2.06 (1.96–2.17)
Female sex	1.24 (1.22–1.26)	1.17 (1.15–1.19)
Province/territory		
Newfoundland and Labrador	1.08 (1.03–1.13)	1.01 (0.97–1.06)
Prince Edward Island	1.32 (1.26–1.39)	1.12 (1.05–1.18)
Nova Scotia	1.87 (1.81–1.93)	1.60 (1.55–1.66)
New Brunswick	1.45 (1.41–1.51)	1.33 (1.28–1.38)
Quebec	1.00 (ref)	1.00 (ref)
Ontario	1.40 (1.37–1.43)	1.21 (1.16–1.26)
Manitoba	1.22 (1.17–1.27)	1.13 (1.08–1.18)
Saskatchewan	1.26 (1.21–1.30)	1.16 (1.12–1.21)
Alberta	1.25 (1.21–1.29)	1.19 (1.15–1.24)
British Columbia	1.32 (1.28–1.36)	1.25 (1.21–1.29)
Nunavut	1.55 (1.45–1.67)	2.54 (2.32–2.78)
Northwest Territories	1.49 (1.41–1.58)	1.88 (1.76–2.02)
Yukon	1.46 (1.29–1.45)	1.36 (1.27–1.45)
Chronic condition		
Asthma	1.29 (1.26–1.32)	1.24 (1.21–1.27)
COPD	1.64 (1.59–1.69)	1.20 (1.16–1.23)
Heart disease	2.13 (2.09–2.18)	1.20 (1.17–1.22)
Stroke	1.87 (1.79–1.95)	1.02 (0.98–1.06)
Diabetes	2.05 (2.01–2.09)	1.25 (1.23–1.28)
Cancer	1.92 (1.86–1.98)	1.15 (1.12–1.19)
Has regular doctor	2.52 (2.43–2.61)	1.67 (1.61–1.74)
Influenza season		
2006/07	1.00 (ref)	1.00 (ref)
2007/08	0.93 (0.90–0.96)	0.93 (0.91–0.96)
2008/09	0.94 (0.91–0.97)	0.92 (0.90–0.95)
2009/10	0.81 (0.78–0.83)	0.91 (0.87–0.94)
2010/11	0.89 (0.86–0.92)	0.86 (0.84–0.89)
2011/12	0.87 (0.85–0.90)	0.81 (0.78–0.84)
2012/13	0.87 (0.84–0.90)	0.79 (0.76–0.82)
2013/14	0.96 (0.93–0.99)	0.85 (0.82–0.88)

Note: CI = confidence interval, COPD = chronic obstructive pulmonary disease, ref = reference category. \*The multivariable regression model included a sample of 456 881 respondents for whom there was complete information on all covariates. In addition to the variables in this table, the model was adjusted for the following: location of residence, household income, educational attainment, marital status, immigration status, racial/cultural group, self-reported health status, category of body mass index, having a child less than 5 yr of age in household, smoking status and receipt of influenza A/H1N1 pandemic vaccine in 2009 (see Appendix 2, available at [www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1](http://www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.151027/-/DC1)).

be associated with influenza vaccination. It also allowed us to include both pre- and post-policy data from provinces that implemented pharmacist policies.

The major limitation of our study was the unavailability of information on the vaccine provider. We were therefore unable to determine the extent of provider “displacement” (i.e., individuals who previously received the vaccine from a physician or nurse and were now getting it at a pharmacy).

We were also unable to measure the extent to which pharmacists implemented the practice of providing influenza vaccines in provinces that implemented a pharmacist policy. This may have varied by province and duration. The absence of a before–after effect in certain provinces may have been due to underlying temporal trends (e.g., a decrease in coverage over time that may have been mitigated by introduction of a pharmacist policy) or incomplete penetrance of the policy (e.g., in Ontario, not all pharmacists complete the training to administer vaccines and not all pharmacies undergo the approval process that is required to provide influenza vaccines).

Another limitation is the Canadian Community Health Survey's exclusion of important populations, such as children less than 12 years of age and people in institutions, both of which are important groups for influenza vaccination. However, because Canadian pharmacists are not authorized to administer influenza vaccines to children (< 5 yr of age in some provinces and < 9 yr of age in others) and institutions often have established vaccination programs, these groups may be less likely to obtain the vaccine from pharmacists.

We were also limited by the data available from the Canadian Community Health Survey. Data from the most recent survey period, during which additional provinces implemented a pharmacist policy, are not yet available.

Finally, our study relied on self-report of influenza vaccination status within the 12 months before survey participation, which may be susceptible to reporting biases. However, self-report of influenza vaccination has been shown to be valid<sup>36–39</sup> and has been used previously to estimate vaccine coverage.<sup>23,40,41</sup>

## Conclusion

The presence of a policy allowing pharmacists to administer publicly funded seasonal influenza vaccines was associated with modest increases in vaccine uptake. Future work to characterize the populations that access this service will aid evaluation efforts and inform decision-making in other jurisdictions that are considering expanding the scope of pharmacist practice to include delivery of vaccines and other services. Ongoing efforts will be needed to evaluate the longer-term impact of pharmacist policies.

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